

where $H(\xi)$ is the entropy of the random variable ξ . The inequality (10) becomes an equation if and only if ξ is uniformly distributed on $B_{n,t}$. If ξ is uniformly distributed, the sequence of random variables ξ_1, \dots, ξ_n forms a Markov chain of complexity $\tau - 1$. Then

$$\ln |A_{n,t}| \leq I(\eta, \eta') = H(\eta) - H(\xi) = Cn, \quad (11)$$

where η is a random variable with a uniform distribution on a set of sequences $\bar{y} \in Y^n$, independent of ξ , $\eta' = \eta + \xi$, $H(\eta)$ is the amount of information in a pair (η, η') , C is the capacity of a channel with noise ξ . Formula (11) may be obtained as follows. We denote by $\bar{\eta}$ a random variable with a uniform distribution on the set $A_{n,t}$; then the values $\bar{\eta}$ are uniquely defined by the values of the random variable $\bar{\eta} = \bar{\eta} + \xi$, hence $\ln |A_{n,t}| = H(\bar{\eta}) = I(\bar{\eta}, \bar{\eta}')$; on the other hand $I(\bar{\eta}, \bar{\eta}') \leq I(\eta, \eta') = Cn$. Determining an exact expression for the entropy $H(\xi)$ is an extremely difficult problem. However, it happens that an estimate of the entropy of the simpler Markov chains enables fairly good estimates of $|B_{n,t}|$ to be obtained.

We consider a Markov chain $\dots, \xi_{l-1}, \xi_l, \xi_{l+1}, \dots$ of complexity $\tau - 1$ with transition probabilities

$$p(z_l) = \begin{cases} p_{0l} & \text{if } \mu(z_{l-1}, \dots, z_{l-\tau+1}) < t, \\ 0 & \text{if } \mu(z_{l-1}, \dots, z_{l-\tau+1}) = t, \end{cases} \quad (12)$$

defined on infinite two-sided sequences $(\dots, z_{-1}, z_0, z_1, \dots)$. This Markov chain is homogeneous and ergodic and a stationary distribution of states $p(z_1, \dots, z_{\tau-1})$ is defined for it, and consequently, a stationary Markov chain with transition probabilities (12) exists. We denote by H its entropy per step. We define the random variable $\xi = (\xi_1, \dots, \xi_n)$ as a segment of this Markov chain of length n . We have

$$H(\xi) \geq nH \quad (13)$$

and

$$\begin{aligned} \frac{1}{n} \ln |B_{n,t}| &\geq H = \\ &= - \sum_{z_1, \dots, z_{\tau-1}} \sum_{\mu(z_1, \dots, z_{\tau-1}) < t} p(z_1, \dots, z_{\tau-1}) \ln p(z_1, \dots, z_{\tau-1}) = \\ &= - \sum_{\mu(z_1, \dots, z_{\tau-1}) < t} p(z_1, \dots, z_{\tau-1}) [(q-1)p_0 \ln p_0 + (1-(q-1)p_0) \times \\ &\quad \times \ln(1-(q-1)p_0)] = \sum_{\mu(z_1, \dots, z_{\tau-1}) < t} p(z_1, \dots, z_{\tau-1}) [H((q-1)p_0) + \\ &\quad + (q-1)p_0 \ln(q-1)]. \end{aligned} \quad (14)$$

To estimate H we need to estimate the probability

$$\sum_{\mu(z_1, \dots, z_{\tau-1}) < t} p(z_1, \dots, z_{\tau-1}) = \bar{p}_{st}.$$

We put $\mu(z_1, \dots, z_{\tau-1}) = t$ and consider the state $(z_{\tau+s}, \dots, z_{\tau+s-1})$, $s \geq 0$, of a Markov chain such that $\mu(z_1, \dots, z_{\tau-1}) = \mu(z_{\tau+s}, \dots, z_{\tau+s-1}) = t$ and $\mu(z_{\tau-l}, \dots, z_{\tau-l-1}) < t$, $0 \leq l < s$, and let \bar{t} be the mean value of s , and \bar{m}_t the mean number of states of weight less than t between the states $(z_1, \dots, z_{\tau-1})$ and $(z_{\tau+s}, \dots, z_{\tau+s-1})$. We have $\bar{p}_{st} = \frac{\bar{m}_t}{\tau + \bar{s} - 1}$. Also we denote by \bar{m} the mean number of zeros between two successive nonzero symbols of the Bernoulli scheme ξ_1, ξ_2, \dots with $p(z) = p_0$, $z = 1, q-1$, that is,

$$\bar{m} = \frac{1-(q-1)p_0}{(q-1)p_0}.$$

We shall show that

$$\bar{m}_t \geq \bar{m}t. \quad (15)$$

To prove (15) we notice that the Markov chain with the transition probabilities (12) can be obtained from the one-sided Markov chain $\dots, \xi_{-2}, \xi_{-1}, \xi_0$ and the sequence of random variables ξ_1, ξ_2, \dots as follows:

the code A corrects t errors if a method of decoding the code $\alpha(\bar{y}) = \bar{y} \in A$, $\bar{y} \in Y^n$, with delay τ exists such that

$$\bar{y} = \bar{\alpha}(\bar{y} + \bar{z}) \quad (3)$$

for $\bar{y} \in A$ and any sequence \bar{z} with $\mu(\bar{z}) \leq t$.

In the case of block codes with block length τ the weight of the sequence \bar{z} is defined as

$$\mu_{bl}(\bar{z}) = \max \mu(z_{i+d}, \dots, z_{i+d+\tau-1}), \quad l = 0, \frac{n}{\tau} - 1, \quad (4)$$

and the code distance $\rho_{bl}(A)$ is defined by the relation

$$\rho_{bl}(A) = \min_{\bar{y}, \bar{y}' \in A, \bar{y} \neq \bar{y}'} \rho(\bar{y}, \bar{y}'),$$

where $\rho(\bar{y}, \bar{y}') = 0$ is the number of nonzero symbols of the sequence $\bar{y}'_{i+d} - \bar{y}_{i+d}, \dots, \bar{y}'_{i+d+\tau-1} - \bar{y}_{i+d+\tau-1}$, and $\bar{y}'_i = \bar{y}_i$ if $i \neq i+d$.

In the case of tree codes with arc length τ if

$$\mu_{tr}(\bar{z}) = \max \mu(z_{i+d}, \dots, z_{i+d+\tau-1}), \quad s = \frac{\tau}{\tau_0}, \quad l = 0, \frac{n}{\tau_0} - s, \quad (5)$$

if $s = 1$ the tree code is a block code.

It is obvious that

$$\mu(\bar{z}) \geq \mu_{tr}(\bar{z}) \geq \mu_{bl}(\bar{z}), \quad \rho_{bl}(A) \leq \rho(A). \quad (6)$$

We will only stop to consider the weights $\mu(\bar{z})$ and $\mu_{bl}(\bar{z})$.

We introduce some notation: $A_{n,t}$, $A_{bl,n,t} \subseteq Y^n$ are respectively the codes of greatest length with code distance $\mu(A_{n,t}) \geq 2t + 1$, $\rho_{bl}(A_{bl,n,t}) \geq 2t + 1$; $B_{n,t}$ and $B_{bl,n,t}$ are the sets of all the sequences \bar{z} of length n of weight $\mu(\bar{z}) \leq t$ and $\mu_{bl}(\bar{z}) \leq t$. By reason of the inequalities (6) we have inequalities for the number of elements of the sets.

$$|B_{bl,n,t}| \geq |B_{n,t}|, \quad |A_{bl,n,t}| \leq |A_{n,t}|.$$

§3. Upper Bounds for the Size of the Code

Theorem 1. For the code A to correct t errors, it is necessary and sufficient that $\rho(A) \geq 2t + 1$.

The proof is obvious. By Theorem 1

$$|A_{n,t}| \leq \frac{q^n}{|B_{n,t}|}.$$

We have

$$|B_{bl,n,t}| = \sum_{i=0}^t \binom{n}{i} (q-1)^{n-i} \quad (8)$$

and

$$|A_{bl,n,t}| \leq \frac{q^n}{\left(\sum_{i=0}^t (q-1)^i C_i^t \right)^{n/t}}. \quad (9)$$

Formulas (8) and (9) are well known in the case where $\tau = n$ (the transmission of one block).

Formula (9) is called Hamming's bound for the code length.

We now estimate the powers of the sets $|B_{n,t}|$ and $|A_{n,t}|$. For this we consider the random variable $\xi = (\xi_1, \dots, \xi_n)$ with values on the set $B_{n,t}$. It is easy to see that

$$\ln |B_{n,t}| \geq H(\xi),$$

Everywhere in this paper where nothing is said to the contrary, logarithms to the natural base are considered.

Theorem 2.

$$|B_{n,t}| = \left[\frac{\tau^t (q-1)^t}{\ln^t \tau} \dots (1+o(1)) \right]^{n/t}, t = \text{const}, n \rightarrow \infty, \tau \rightarrow \infty. \quad (20)$$

It is obvious that the expressions in brackets on the right sides of (8) and (20) differ by the factor $\frac{t!}{\ln^t \tau} (1+o(1))$.

The problem of constructing the code length

$$q^n \left[\frac{d \tau^t (q-1)^t}{\ln^t \tau} \right]^{n/t}, d = \text{const} \quad (16)$$

is extremely interesting. For $t=1$ and $q=2$, Ash and Wyner [6] constructed convolutional tree codes of length

$$|A_{n,t}| = \left(\frac{d 2^t}{\tau / \log_2 \tau} \right)^{n/t}.$$

and consequently,

$$\begin{aligned} \bar{m}_t &\geq \max \left(\frac{t[1-(q-1)p_0]}{(q-1)p_0}, \bar{s} \right), \\ \bar{p}_{s,t} &= \frac{\bar{m}_t}{\tau + \bar{s} - 1} \geq \max \left[\frac{t[1-(q-1)p_0]}{(q-1)p_0}, \frac{\bar{s}}{\tau + \bar{s}} \right] \geq \\ &\geq \frac{t(1-(q-1)p_0)}{(q-1)p_0} = \frac{t(1-(q-1)p_0)}{\tau(q-1)p_0 + t(1-(q-1)p_0)}. \end{aligned}$$

Putting $p_0 = \frac{\ln \tau}{(q-1)\tau}$, from the last relation and from (14) we obtain

$$\begin{aligned} &\frac{1}{n} \ln |B_{n,t}| \geq \bar{H} \geq \\ &\geq -\frac{\ln \tau}{\tau} + t \left(1 - \frac{\ln \tau}{\tau} \right) \left[\frac{\ln \tau}{\tau} \ln \frac{\ln \tau}{(q-1)\tau} + \left(1 - \frac{\ln \tau}{\tau} \right) \ln \left(1 - \frac{\ln \tau}{\tau} \right) \right]. \end{aligned}$$

From this as $\tau \rightarrow \infty$ we obtain

$$|B_{n,t}| \geq \left[\frac{\tau^t (q-1)^t}{\ln^t \tau} (1+o(1)) \right]^{n/t}.$$

We now prove the converse inequality

$$|B_{n,t}| \leq \left[\frac{\tau^t (q-1)^t}{\ln^t \tau} (1+o(1)) \right]^{n/t}.$$

In the case $t=1$, $q=2$ an expression for $|B_{n,t}|$ was found by Vershinski and Mikhelev [5]

$$|B_{n,t}| = \sum_{i=0}^t C_{n-t-i}^{i-1} C_a^b = 0 \quad \text{if } b > a.$$

This relation is easily generalized to the case $q > 2$:

$$|B_{n,t}| = \sum_{i=0}^t C_{n-t-i}^{i-1} (q-1)^i = \left[\frac{\tau^t (q-1)^t}{\ln^t \tau} (1+o(1)) \right]^{n/t}, n, \tau \rightarrow \infty. \quad (17)$$

The inequality (18) follows from the relation (19) and the inequality

$$|B_{n,t}| \leq |B_{n,t-1}|.$$

In order to obtain the last inequality, it is sufficient to notice that if in the sequence of the set $B_{n,t}$ the i -th zero symbols, $s=1, 2, \dots$, are replaced by zeros, the sequence of the set $B_{n,t-1}$ is obtained, but if only the i -th zero remains nonzero, the sequence of $B_{n,t}$ is obtained. It is also easy to show that

$$|B_{n,t}| \leq |B_{n,t-1}|.$$

Therefore we have the following theorem.

where $\zeta_i, i=1, \tau-1$ are independent binary random variables, and $p(\zeta_i=1) = (q-1)p_0$. The inequality (21) obviously follows from the proof of the inequality (15). In order to select the value of p_0 minimizing the right side of (14), we use the estimate (see, for example, [12])

$$1 - \bar{p}_{s,t} \leq p(\zeta_1' + \dots + \zeta_{\tau-1}' \geq t), \quad (21)$$

$$p(\zeta_1' + \dots + \zeta_{\tau-1}' + (q-1)p_0(\tau-1) \geq x\sigma\sqrt{\tau}) \leq \frac{e^{-x}}{\sqrt{\tau}} + \Phi(x)$$

$$\varepsilon_1 \leq \frac{1.33M(\zeta_1' - (q-1)p_0)^3}{\sigma^3} \leq 1.33/\sigma; \quad \sigma^2 = (q-1)p_0(1-(q-1)p_0);$$

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_0^x e^{-x^2/2} dx. \quad (22)$$

Putting $(q-1)p_0 = \frac{t}{\tau} - \delta$, where $\delta = \sqrt{\frac{\sigma^2 \ln \tau}{\tau}}$ we obtain from (22)

$$\begin{aligned} 1 - \bar{p}_{s,t} &\leq p(\zeta_1' + \dots + \zeta_{\tau-1}' \geq t) = \\ &= p\left\{ \zeta_1' + \dots + \zeta_{\tau-1}' - (q-1)p_0\tau \geq \frac{\delta\sqrt{\tau}}{\sigma} \sigma\sqrt{\tau} \right\} \leq \\ &\leq \frac{e^{-\frac{\delta\sqrt{\tau}}{\sigma}}}{\sqrt{\tau}} + \frac{1}{\delta\sigma\sqrt{2\pi\tau}} e^{-\left(\frac{\delta\sqrt{\tau}}{\sigma}\right)^2/2} = \frac{e_1}{\sqrt{\tau}} + \frac{1}{\sqrt{2\pi\tau} \ln \tau}. \end{aligned} \quad (23)$$

Substituting $\frac{t}{\tau} - \delta = (q-1)p_0$ in (14) and using the inequality $H\left(\frac{t}{\tau} - \delta\right) \geq H\left(\frac{t}{\tau}\right) + \delta \ln \frac{\tau-t+\tau\delta}{t-\tau\delta}$, $\delta > 0$,

and the relation (23), we obtain:

Theorem 3.

$$\begin{aligned} \frac{1}{n} \ln |B_{n,t}| &\geq H\left(\frac{t}{\tau}\right) + \frac{t}{\tau} \ln(q-1) + (1+o(1)) \sqrt{\frac{\sigma^2 \ln \tau}{\tau}} \times \\ &\times \left[\ln \frac{\tau-t}{1-t} - \ln(q-1) \right], \end{aligned} \quad (24)$$

where $H(p_0) = -p_0 \ln p_0 - (1-p_0) \ln(1-p_0)$, $t = \alpha\tau$, $\alpha < 1$, $n \rightarrow \infty$, $\tau \rightarrow \infty$.

For $q=2$, formula (24) assumes the form

$$\frac{1}{n} \ln |B_{n,r}| \geq H\left(\frac{t}{\tau}\right) + (1 + o(1)) \sqrt{\frac{\sigma^2 \ln \tau}{\tau}} \ln \frac{t}{1 - \frac{t}{\tau}}.$$

This expression differs from the corresponding expression for a block code by the presence of the second term of order

$$o\left(\sqrt{\frac{\ln \tau}{\tau}}\right).$$

Theorem 4. (Elias bound, see [8, 9]). For binary codes

$$R = \frac{1}{n} \ln |A_{n,r}| \leq \ln 2 - H\left(\frac{1 - \sqrt{1 - 4 \frac{t}{\tau}}}{2}\right) + o\left(\sqrt{\frac{\ln \tau}{\tau}}\right).$$

We prove the theorem by means of the following lemmas.

Let D be a set of sequences of n symbols with the j -initial symbols in common; we call the number of branches of the set D on the segment $[j+1, j+l]$, the number of different subsequences of the sequences of D on the segment $[j+1, j+l]$.

Lemma 1. If $(r+l)^2 > \frac{1}{2}(2(r+l) - d)(l + \tau)$ the subset D of the set $B_{n,r}$ of sequences with code distance $\rho(D) \geq d = 2r + 1$ and j common initial symbols, has not more than

$$\frac{1/2d(l + \tau)(r + l)}{(r + l)^2 - 1/2(2(r + l) - d)(l + \tau)}$$

branches on the segment $[j+1, j+l]$.

The proof of Lemma 1 is similar to the proof of the lemma in [9]. We denote by $|A_{n,r}|$ the maximal number of elements of the subset of the set $B_{n,r}$ with code distance $d = 2r + 1$. It follows from Lemma 1 that

$$|A_{n,r}| \leq \left[\frac{1/2d(l + \tau)(r + l)}{(r + l)^2 - 1/2(2(r + l) - d)(l + \tau)} \right]^{1/n}.$$

Lemma 2.

$$|A_{n,r}| |B_{n,r}| \leq 2^n |A_{n,r,r}|.$$

This formula is similar to formula (5) of [9]. Putting

$$r + l = \frac{1}{2}(\tau + l) \left(1 - \sqrt{1 - \frac{4l}{\tau}}\right), \quad l = \sqrt{\tau \ln \tau},$$

we obtain from (28)

$$|A_{n,r}| \leq \frac{2^n (d(r + \sqrt{\tau \ln \tau}))}{\sqrt{\tau \ln \tau}}.$$

Comparing (29) and (31), we find that

$$|A_{n,r}| \leq \frac{2^n (d(r + \sqrt{\tau \ln \tau}))}{\sqrt{\tau \ln \tau}} |B_{n,r}|.$$

By (25),

$$\frac{1}{n} \ln |B_{n,r}| \geq H\left(\frac{r}{\tau}\right) + o\left(\sqrt{\frac{\ln \tau}{\tau}}\right) = H\left(\frac{1 - \sqrt{1 - 4 \frac{t}{\tau}}}{2}\right) + o\left(\sqrt{\frac{\ln \tau}{\tau}}\right).$$

Result (24) follows from (33) and (32). The theorem is proven.

Estimation of the Error Probability

In investigating the error probability we restrict ourselves to the case of a binary symmetric memoryless channel; we consider the message to be given as a sequence of equidistributed independent binary random variables: $\xi = (\xi_1, \dots, \xi_m)$, $P_{\xi_i}(1) = P_{\xi_i}(0) = 1/2$, $i = \overline{1, m}$. The transmission rate along the channel is $R = \frac{m}{n} \ln 2$.

Encoding defines the random variable $\eta = f(\xi)$, the input signal; the transition probability of the channel p_0 defines the random variable η' , the output signal, such that

$$P_{\eta\eta'}(\bar{y}, \bar{y}') = P_{\eta\eta'}(y_1, \dots, y_n; y'_1, \dots, y'_n) = P_{\eta_1}(y_1, \dots, y_n) \times \\ \times P(y'_1, \dots, y'_n | y_1, \dots, y_n) = P_{\eta_1}(y_1, \dots, y_n) \prod_{i=1}^n P(y'_i | y_i),$$

finally, decoding defines the random variable $\xi' = (\xi'_1, \dots, \xi'_m) = \varphi(\eta')$. The sequence ξ , η , η' , ξ' forms a Markov chain.

We define the error probability in the transmission (f, φ) as the number

$$P_e(f, \varphi) = \frac{1}{m} \sum_{i=1}^m P(\xi'_i \neq \xi_i) = \frac{1}{m} \sum_{i=1}^m \sum_{x_i \neq x'_i} P_{\xi_i} \varphi_i(x_i, x'_i). \quad (34)$$

If (f, φ) is a block transmission with decoding delay τ , we know [7] the lower and upper limits of the lower bound $P_{e, \text{opt}}(R) = \inf P_e(f, \varphi)$ of the error probability $P_e(f, \varphi)$, which are identical apart from the factor for (f, φ)

$$C \geq R \geq R_{\text{crit}} = 1 - H\left(\frac{\sqrt{p_0}}{\sqrt{p_0} + \sqrt{1 - p_0}}\right). \quad (35)$$

The question arises of whether it is possible to decrease the error probability by using a nonblock method of transmission. The answer to this question is given by the following theorem.

Theorem 5. In transmission with delay τ along a binary symmetric memoryless channel

$$P_{e, \text{opt}}(R) \equiv P_{e, \text{opt}}(R, \tau) \geq \\ \geq \frac{p_0(1 - \rho)}{4(1 - p_0)\sqrt{2\tau\rho(1 - \rho)}} e^{-\tau \left(\frac{\rho(1 - \rho)}{p_0} + (1 - \rho) \ln \frac{1 - \rho}{1 - p_0} \right)} \equiv P_e(R, \tau) \quad (36)$$

for

$$R = \frac{m}{n} \ln 2 \geq \frac{\ln 2 - H(\rho) - \sqrt{\rho(1 - \rho)} \ln \tau}{1 - \left[H\left(\frac{\varepsilon_1}{\sqrt{\tau}}\right) - \frac{1}{\sqrt{2\tau \ln \tau}} \ln \frac{\varepsilon_1}{\sqrt{\tau - \varepsilon_1}} \right] \frac{1}{\sqrt{\tau - \varepsilon_1}}}, \quad \varepsilon_1 \leq 1.33/\sigma \quad (37)$$

where ρ is a parameter, $p_0 + \sqrt{\frac{\sigma^2 \ln \tau}{\tau}} \leq \rho \leq \frac{1}{2}$, $\sigma^2 = \sqrt{p_0(1 - p_0)}$.

Proof. It has to be proved that for any method of transmission with decoding delay

$$P_e(f, \varphi) = \frac{1}{m} \sum_{i=1}^m P(\xi'_i \neq \xi_i) \geq P_e(R, \tau), \quad (38)$$

where $\xi'_i = \varphi_i(\eta'_1, \dots, \eta'_{j+\tau-1})$, $\eta'_j = y'_{j,j}$ is the initial symbol of the i -th symbol of the x_i -message. The essence of the proof reduces to the following. From the value of the error probability $P_e(f, \varphi)$ for the transition probability p_0 of the channel, the quantity of information $I(\xi, \eta')$ is estimated for the transmission probability $\tilde{p}_0 \geq p_0$ and for the same transition method (f, φ) . The inequality (36) is obtained from the condition that $\frac{1}{n} I(\xi, \eta')$ cannot be greater than the transmission capacity of the channel $C(p_0)$ for a transition probability $\tilde{p}_0 \geq p_0$. We will now deduce the correspond-

ing bounds. Since the noise $\tilde{\xi} = (\xi_1, \dots, \xi_m) = (\eta_1' - \eta_1, \dots, \eta_m' - \eta_m)$ is independent of the message $\tilde{\xi} = (\xi_1, \dots, \xi_m)$, the sequence of random variables

$$\xi_i, (\eta_1, \dots, \eta_{i-1}), (\eta_1', \dots, \eta_{i-1}'), (\eta_1, \dots, \eta_{i-1}, \eta_{i-1}', \dots, \eta_{i-1}, \eta_{i-1}') \quad (39)$$

forms a Markov chain, and therefore on decoding for the maximum a posteriori probability defined by the expression

$$(40)$$

$$\tilde{\xi}_i' = \tilde{\varphi}_i(\eta_1, \dots, \eta_{i-1}, \eta_{i-1}', \dots, \eta_{i-1}, \eta_{i-1}'),$$

where

$$P\{\xi_i = \tilde{\xi}_i' / \eta_1, \dots, \eta_{i-1}, \eta_{i-1}', \dots, \eta_{i-1}, \eta_{i-1}'\} \geq P\{\xi_i \neq \tilde{\xi}_i' / \eta_1, \dots, \eta_{i-1}, \eta_{i-1}', \dots, \eta_{i-1}, \eta_{i-1}'\} \quad (41)$$

$$\geq P\{\xi_i \neq \tilde{\xi}_i' / \eta_1, \dots, \eta_{i-1}, \eta_{i-1}', \dots, \eta_{i-1}, \eta_{i-1}'\} \quad (42)$$

we have

$$P\{\xi_i \neq \tilde{\xi}_i'\} \geq P\{\xi_i \neq \tilde{\xi}_i'\}.$$

Also we denote by $A_i(\eta_1, \dots, \eta_{i-1}, \eta_{i-1}', \dots, \eta_{i-1}, \eta_{i-1}')$ the set of all subsequences Z_{ij}, \dots, Z_{i+j-1} such that if

$$\eta_1 = y_1, \dots, \eta_{i+j-1} = y_{i+j-1}, \eta_1' = y_{i+j} + z_{i+j}, \dots, \eta_{i+j-1}' = y_{i+j-1} + z_{i+j-1} + z_{i+j},$$

then $\xi_i = \tilde{\xi}_i'$. We have

$$P\{\xi_i \neq \tilde{\xi}_i'\} = 1 - \sum_{(y_1, \dots, y_{i+j-1}) \in A_i} P_A(y_1, \dots, y_{i+j-1}, y_{i+j}, \dots, y_{i+j-1}) \times \prod_{(z_1, \dots, z_{i+j-1}) \in A_i(y_1, \dots, y_{i+j-1})} \frac{P_{\eta_1}(y_1, \dots, y_{i+j-1}, y_{i+j}, \dots, y_{i+j-1})}{P_{\eta_1}(y_1, \dots, y_{i+j-1})} \quad (43)$$

and for the transition probability \tilde{p}_0

$$\begin{aligned} & \sum_{(z_1, \dots, z_{i+j-1}) \in A_i(y_1, \dots, y_{i+j-1})} \frac{P_{\eta_1}(y_1, \dots, y_{i+j-1}, y_{i+j}, \dots, y_{i+j-1})}{P_{\eta_1}(y_1, \dots, y_{i+j-1})} = \\ & = \sum_{(z_1, \dots, z_{i+j-1}) \in A_i(y_1, \dots, y_{i+j-1})} \frac{1}{\tilde{p}_0} \sim \frac{1}{\tilde{p}_0} (z_1, \dots, z_{i+j-1}) \quad (1 - \tilde{p}_0)^{-(i+j-1)} \quad (44) \end{aligned}$$

It is easy to verify that the right side of (44) is majorized by the sum over the spherical domain $A_{sp}(y_1, \dots, y_{i+j-1})$ containing all possible sequences (z_1, \dots, z_{i+j-1}) of weight $s \leq i$, and not containing sequences of weight $t + 2$, and that if $\tilde{p}_0 = p_0$ this sum equals the right side of (44). Similarly (43) is majorized by the sum over the sets (spheres) $A_{sp}(y_1, \dots, y_{i+j-1})$ containing all possible sequences (z_1, \dots, z_{i+j-1}) of weight t and not containing sequences of weight $t + 2$, where t is independent of the code sequence (y_1, \dots, y_n) and of the subscript $i = \overline{1, n}$.

Therefore the inequality

$$\begin{aligned} \tilde{p}_0 &= \frac{1}{m} \sum_{i=1}^m P\{\xi_i \neq \tilde{\xi}_i'\} \leq \\ & \leq 1 - \sum_{i=0}^t C_t \tilde{p}_0^i (1 - \tilde{p}_0)^{t-i} = \sum_{i=t+1}^m C_t \tilde{p}_0^i (1 - \tilde{p}_0)^{t-i} \quad (45) \end{aligned}$$

is valid for any values $\tilde{p}_0 \geq p_0$, if it is valid for $\tilde{p}_0 = p_0$.

We estimate the quantity of information $I(\xi, \eta')$ if $\tilde{p}_0 \geq p_0$. For this we consider the random variable $\beta = (\beta_1, \dots, \beta_m)$

$$\beta_i = \begin{cases} 1, & \text{if } \xi_i' \neq \xi_i, \\ 0, & \text{if } \xi_i' = \xi_i. \end{cases} \quad (46)$$

Obviously,

$$\frac{1}{m} \sum_{i=1}^m P\{\beta_i = 1\} = \frac{1}{m} \sum_{i=1}^m P\{\xi_i \neq \xi_i'\} = \tilde{p}_0, \quad H(\beta) \leq mH(\tilde{p}_0). \quad (47)$$

It is easy to verify that the random variable ξ is a function of the pair of random variables (η', β) . This implies that

$$m \ln 2 = H(\xi) = I(\xi, (\eta', \beta)) = I(\xi, \eta') + M'(\xi, \beta | \eta') \leq I(\xi, \eta') + H(\beta). \quad (48)$$

By the definition of the transmission capacity of the channel

$$I(\xi, \eta') \leq I(\eta', \eta') \leq nC(\tilde{p}_0) = n(1 - H(\tilde{p}_0)). \quad (49)$$

Comparing (47)-(49), we find

$$m \ln 2 - mH(\tilde{p}_0) \leq nC(\tilde{p}_0). \quad (50)$$

We put that $\tilde{p}_0 = \frac{t}{n} - \sqrt{\frac{\sigma^2 \ln 2}{t}}$, $\sigma^2 = \tilde{p}_0(1 - \tilde{p}_0)$, and ξ_1, \dots, ξ_n are independent binary random variables with

$P\{\xi_i = 1\} = \tilde{p}_0$, $i = \overline{1, n}$. Using the inequality (23), we obtain

$$\begin{aligned} \tilde{p}_0 & \leq \sum_{i=1}^t C_t \tilde{p}_0^i (1 - \tilde{p}_0)^{t-i} = P\left\{\sum_{i=1}^t \xi_i > t\right\} \leq \\ & \leq \frac{e^t}{\sqrt{2\pi t} \ln t}. \end{aligned}$$

Moreover, from this

$$H(\tilde{p}_0) \leq H\left(\frac{e^t}{\sqrt{2\pi t} \ln t}\right) - \frac{1}{\sqrt{2\pi t} \ln t} \ln \frac{e^t}{\sqrt{2\pi t} \ln t} \quad (51)$$

$$C(\tilde{p}_0) = \ln 2 - H(\tilde{p}_0) \leq \ln 2 - H\left(\frac{t}{n}\right) - \sqrt{\frac{\sigma^2 \ln 2}{t}} \ln \frac{p_0}{1 - p_0}. \quad (52)$$

Putting $p = t/n$, we deduce from (50)-(52) that if the error probability

$$\begin{aligned} P_e(U, \varphi) &= 1 - \sum_{i=1}^t C_t \tilde{p}_0^i (1 - \tilde{p}_0)^{t-i} \geq \\ & \geq \frac{p_0(1-p)}{4(1-p_0)p} \sqrt{\frac{2\pi t p(1-p)}{t}} e^{-t \left[p \ln \frac{p}{p_0} + (1-p) \ln \frac{1-p_0}{1-p} \right]}, \end{aligned} \quad (53)$$

the transmission rate must satisfy the relation

$$\begin{aligned} R = \frac{m}{n} \ln 2 & \leq \frac{C(\tilde{p}_0)}{1 - H(\tilde{p}_0)} \leq \\ & \leq \frac{\ln 2}{\ln 2 - H(p) - \sqrt{\frac{p(1-p) \ln 2}{t} \ln \frac{p_0}{1-p_0}}} \cdot \frac{1}{1 - \left[H\left(\frac{e^t}{\sqrt{2\pi t} \ln t}\right) - \frac{1}{\sqrt{2\pi t} \ln t} \ln \frac{e^t}{\sqrt{2\pi t} \ln t} \right]} \quad (54) \end{aligned}$$

The theorem is proven.

Remark. The expression (37) for the transmission rate differs from the corresponding expression in [7] (formula (1.50)) by the presence of a term of order $\sqrt{\frac{\ln^2 t}{t}}$.

However, an exact bound has still not been found for the error probability of the random method of constructing

19 April 1967

nonblock codes (in particular, convolutional codes). It remains not clear whether the presence of the term $O(\sqrt{\frac{\ln^2 \tau}{\tau}})$ is caused by an imperfection of the method of obtaining the estimate, or whether this term reflects an improvement of the asymptotic error probability for nonblock transmission as $R \rightarrow C$.

As $R \rightarrow 0$ the asymptotic error probability for nonblock transmission was found in [10, 11]; it happens that the exponential term of this asymptotic function is also retained in nonblock transmission; the latter fact is implied by the following theorem.

Theorem 6. In transmission over a binary symmetric memoryless channel with decoding delay and rate

$$P_s(f, \varphi) \geq e^{-\frac{1}{4} \ln \frac{1}{4\mu(1-\mu)} + O(\ln \tau)} \quad (55)$$

55. Transmission with Feedback

In this section we show that many of the results established in the preceding sections are preserved if feedback is present. The definition of transmission with decoding delay in the presence of feedback is the same as the corresponding definition for the case where feedback is absent. Only the concept of coding is changed, namely, instead of $y_j = f(\bar{x})$ the functions $x_j = f_j(\bar{x}, y_1, \dots, y_{j-1})$ are considered, that is, the value of the j -th symbol of the code sequence is a function of the message sent and of the values of the symbols y_1, \dots, y_{j-1} of the output signal.

Theorem 7. In the transmission of messages with decoding delay τ with feedback and with the correction of t errors

$$\frac{1}{n} H(\xi) = \frac{m}{n} \ln 2 \leq \frac{1}{n} \ln \frac{q^n}{|B_{n,t}|} \quad (56)$$

(Formula (56) is similar to formula (7).) To prove the theorem it is sufficient to notice that in the transmission of the same message value $\bar{x} = (x_1, \dots, x_m)$ to different values of "noise" \bar{z} and \bar{z}' there correspond different values of the output signals \bar{y} and \bar{y}' , that is, the set $B_{\bar{x}} = \{y_j\}$, that is, the set $B_{\bar{x}} = \bar{y}$ has power not less than $B_{n,t}$.

Theorem 8. In transmission over a binary symmetric channel with feedback with decoding delay τ the bounds (36) and (37) are preserved.

The proof of Theorem 5 is based on bounds of the transmission capacity of a binary symmetric memoryless channel. The result of Theorem 8 easily follows from the proof of Theorem 5, if use is made of the fact that the presence of feedback does not increase the transmission capacity of a memoryless channel.

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